

# A GENERALIZED CLASS OF JACK-KNIFED ESTIMATOR FOR PRODUCT OF POPULATION MEANS USING TWO AUXILIARY VARIABLES UNDER MEASUREMENT ERRORS

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**Abstract**— In this paper, we have suggested a generalized estimator for the estimation of product of population means in the presence of measurement errors. The expression of bias and mean square error are derived under measurement errors model. Further, we have also proposed a jack-knife version of the estimators based on Quenouille’s method in presence of measurement errors. Finally, some concluding remarks are made clearly demonstrating the importance of the proposed classes.

**Keywords**— Measurement error, Jack-knife estimator, Unbiasedness, Mean square error.

## I. INTRODUCTION

In sampling from finite population, the estimators based on data explicitly assume that the observations are correct measurements on characteristics of study. But unfortunately this ideal situation is not met in practice for a varied of reasons, such as non-response errors, reporting errors and computational errors etc. when the measurement errors are negligible, the statistical inferences based on observed data continues to remain valid. On the contrary when they are not appreciably small or negligible, the inferences may simply be invalid or inaccurate but may also often lead to unexpected, undesirable and unfortunate consequences.

Over the past several decades, statisticians are paying their attention towards the problem of estimation of parameters in the presence of measurement errors. In survey sampling, the properties of data usually presuppose that the observations are the correct measurements on characteristics being studied. However this assumption is not satisfied in many applications and data is contaminated with measurement errors, such as non-response errors, reporting errors and computing errors. These measurement errors make the result invalid, which are meant for no measurement error case. If measurement errors are very small and we can neglect it, then the statistical inference based on data observed continue to remain valid. On the contrary, when they are not appreciably small and negligible, the inferences may not be simply invalid and inaccurate but may often lead to unexpected, undesirable and unfortunate consequences (see Srivastava and Shalabh, 2001). Some important sources of measurement errors in survey data

are discussed in Cochran (1968), Shalabh (1997), and Singh and Karpe (2008, 2010) studied some estimators of population mean under measurement errors.

For a simple random sample of size  $n$ , let  $(y_{1j}, y_{2j}, x_{1j}, x_{2j})$  be the observed values instead of the true values  $(Y_{1j}, Y_{2j}, X_{1j}, X_{2j})$  on the characteristics  $(Y_1, Y_2, X_1, X_2)$  respectively. Let the measurement error be defined as

$$u_{1,i} = y_{1,i} - Y_{1,i} \tag{1.1}$$

$$u_{2,i} = y_{2,i} - Y_{2,i} \tag{1.2}$$

$$v_{1,i} = x_{1,i} - X_{1,i} \tag{1.3}$$

$$v_{2,i} = x_{2,i} - X_{2,i} \tag{1.4}$$

here  $i=1, 2, \dots, n$

which are stochastic in nature and are uncorrelated with mean zero and variances  $\sigma_{u_1}^2, \sigma_{u_2}^2, \sigma_{v_1}^2$  and  $\sigma_{v_2}^2$  respectively. Further,

let population means of  $(Y_1, Y_2, X_1, X_2)$  be  $(\mu_{X_1}, \mu_{Y_1}, \mu_{X_2}, \mu_{Y_2})$  and the population variance of  $(Y_1, Y_2, X_1, X_2)$  be  $(\sigma_{X_1}^2, \sigma_{Y_1}^2, \sigma_{X_2}^2, \sigma_{Y_2}^2)$  respectively.  $\sigma_{X_1 Y_1}, \sigma_{X_2 Y_2}, \sigma_{X_1 Y_2}, \sigma_{X_2 Y_1}$  and  $\sigma_{Y_1 Y_2}$  be the population correlation coefficient between  $(X_1, Y_1), (X_2, Y_1), (X_1, Y_2), (X_2, Y_2), (X_1, X_2)$  and  $(Y_1, Y_2)$  respectively.

**2 PROPOSED CLASS OF GENERALIZED ESTIMATOR**  
The generalized estimator of product of population mean  $\hat{P}_h$  using estimated product of means  $(Y_1, Y_2)$  be  $\hat{P} = \hat{\mu}_{Y_1} \hat{\mu}_{Y_2}$  and  $\mu_1$  and  $\mu_2$  of auxiliary variable  $X_1$  and  $X_2$  respectively is

$$\hat{P}_h = h \left[ \hat{P}, \frac{\bar{x}_1}{\mu_1}, \frac{\bar{x}_2}{\mu_2} \right] \tag{2.1}$$

where  $(\bar{x}_1, \bar{x}_2)$  are the sample means of  $X_1, X_2$  and  $\hat{P} = \hat{y}_1 \hat{y}_2$  is the estimate of  $P = \mu_{Y_1} \mu_{Y_2}$  respectively for a simple random

variable of size  $n$ . Also,  $h(\cdot)$  is the function of  $\frac{\bar{x}_1}{\mu_1}$  and  $\frac{\bar{x}_2}{\mu_2}$ -satisfying

(i).  $h(P,1,1)=P$  (2.2)

(ii). The function is continuous and bounded in the closed interval of real line  $\mathbf{R}$ .

The first and second order partial derivatives of the function are exist and are continuous and bounded in  $\mathbf{R}$ .

Consider the following terms

$$\begin{aligned} \underline{y}_x &= 1 \left[ \sum_{j=1}^n Y_{1,j} \right] \\ &= \frac{1}{n} \left[ \sum_{j=1}^n u_{1,j} + Y_{1,j} - \mu_{Y_1} + \mu_{Y_1} \right] \\ &= \frac{1}{\sqrt{n}} \left[ \frac{W_{u_1} + W_{Y_1}}{\mu_{Y_1}} \right] + \mu_{Y_1} \end{aligned} \tag{2.3}$$

Similarly, we get

$$\underline{y}_2 = \frac{1}{\sqrt{n}} \left[ \frac{W_{u_2} + W_{Y_2}}{\mu_{Y_2}} \right] + \mu_{Y_2} \tag{2.4}$$

$$\underline{x}_1 = \frac{1}{\sqrt{n}} \left[ \frac{W_{v_1} + W_{X_1}}{\mu_{X_1}} \right] + \mu_{X_1} \tag{2.5}$$

and

$$\underline{x}_2 = \frac{1}{\sqrt{n}} \left[ \frac{W_{v_2} + W_{X_2}}{\mu_{X_2}} \right] + \mu_{X_2} \tag{2.6}$$

On expand  $\hat{P}$  using Maclaurin series expansion and ignore the terms of having the powers greater than two, we have

$$\begin{aligned} \hat{P} &= P \left( 1 + \frac{W_{u_1} + W_{Y_1}}{n^{1/2} \mu_{Y_1}} \right) \left( 1 + \frac{W_{u_2} + W_{Y_2}}{n^{1/2} \mu_{Y_2}} \right) \\ &= P \left( 1 + \frac{W_{u_1} + W_{Y_1}}{n^{1/2} \mu_{Y_1}} + \frac{W_{u_2} + W_{Y_2}}{n^{1/2} \mu_{Y_2}} + \frac{(W_{u_1} + W_{Y_1})(W_{u_2} + W_{Y_2})}{n^{1/2} \mu_{Y_1} n^{1/2} \mu_{Y_2}} \right) \end{aligned} \tag{2.7}$$

### 3 BIAS AND MSE OF THE PROPOSED CLASS OF GENERALIZED ESTIMATOR

To obtain the bias of the generalized class of estimator we expand (2.1) by Taylor's series expansion up to third order terms about the point  $(P,1,1)$  and noting that  $h(P,1,1)=P$  and is continuous and bounded in the closed interval of real line  $\mathbf{R}$ . We assume that the first and second order partial derivatives of the function  $h(\cdot)$  are exist and are continuous and bounded

in  $\mathbf{R}$ .

$$\begin{aligned} \hat{P} - P &= |P| \left[ \frac{W_{u_1} + W_{Y_1}}{n^{1/2} \mu_{Y_1}} + \frac{W_{u_2} + W_{Y_2}}{n^{1/2} \mu_{Y_2}} + \frac{(W_{u_1} + W_{Y_1})(W_{u_2} + W_{Y_2})}{n^{1/2} \mu_{Y_1} n^{1/2} \mu_{Y_2}} \right] h \\ &+ \frac{W_{v_1} + W_{X_1}}{n^{1/2} \mu_{X_1}} h_2 + \frac{W_{v_2} + W_{X_2}}{n^{1/2} \mu_{X_2}} h_3 + \frac{1}{2! n} \left\{ P^2 \left[ \frac{W_{u_1} + W_{Y_1}}{\mu_{Y_1}} + \frac{W_{u_2} + W_{Y_2}}{\mu_{Y_2}} \right]^2 \right\} h_{300} \end{aligned}$$

$$\begin{aligned} &+ \frac{(W_{u_1} + W_{Y_1})^2}{\mu_{Y_1}^2} h_{030} + \frac{(W_{u_2} + W_{Y_2})^2}{\mu_{Y_2}^2} h_{003} + 2 \frac{P(W_{u_1} + W_{Y_1})(W_{u_2} + W_{Y_2})}{\mu_{Y_1} \mu_{Y_2}} h_{110} \\ &+ 2 \frac{P(W_{u_1} + W_{Y_1})(W_{v_1} + W_{X_1})}{\mu_{Y_1} \mu_{X_1}} h_{101} + 2 \frac{(W_{v_1} + W_{X_1})(W_{v_2} + W_{X_2})}{\mu_{X_1} \mu_{X_2}} h_{011} \\ &+ 2 \left. \left. \left. \frac{P(W_{u_1} + W_{Y_1})(W_{v_2} + W_{X_2})}{\mu_{Y_1} \mu_{X_2}} h_{101} + 2 \frac{P(W_{u_2} + W_{Y_2})(W_{v_1} + W_{X_1})}{\mu_{Y_2} \mu_{X_1}} h_{011} \right\} \right\} \end{aligned} \tag{3.1}$$

On taking expectation on both side of (3.1), the bias of  $\hat{R}_g$  is

$$\begin{aligned} \text{Bias}(\hat{P}_g) &= \frac{P}{n} \left[ \frac{1}{2\mu^2} \sigma_{X_1}^2 h_{030} + \frac{1}{2\mu^2} \sigma_{X_2}^2 h_{003} + P \mu \mu \sigma_{X_1 X_2} h_{011} \right] \\ &+ \frac{1}{\mu_{Y_1}} \left[ \frac{1}{\mu_{Y_1}} \sigma_{X_1 Y_1} h_{110} + \frac{1}{\mu_{Y_2}} \sigma_{X_2 Y_1} h_{101} \right] + \frac{1}{\mu_{Y_2}} \left[ \frac{1}{\mu_{Y_1}} \sigma_{X_1 Y_2} h_{110} + \frac{1}{\mu_{Y_2}} \sigma_{X_2 Y_2} h_{101} \right] \\ &+ \frac{1}{n} \left[ \frac{1}{2\mu^2} \sigma_{v_1}^2 h_{030} + \frac{1}{2\mu^2} \sigma_{v_2}^2 h_{003} \right] \end{aligned} \tag{3.2}$$

Now squaring the both sides of (3.1) and taking expectation, we get mean Square Error of  $\hat{P}_g$  given by

$$\begin{aligned} \text{MSE}(\hat{P}_g) &= E(\hat{P}_g - P)^2 \\ &= \frac{1}{n} \left[ \frac{1}{\mu_{Y_1}^2} \left( \frac{1}{\mu_{Y_1}^2} \sigma_{Y_1}^2 + \frac{1}{\mu_{Y_2}^2} \sigma_{Y_2}^2 + \frac{2}{\mu_{Y_1} \mu_{Y_2}} \sigma_{Y_1 Y_2} \right) h^2 \right. \\ &+ \left. \left( \frac{1}{P^2 \mu_{X_1}^2} \sigma_{X_1}^2 h^2 + \frac{1}{P^2 \mu_{X_2}^2} \sigma_{X_2}^2 h^2 + \frac{2}{P^2 \mu_{X_1} \mu_{X_2}} \sigma_{X_1 X_2} h h \right) \right. \\ &+ \left. \frac{2}{\mu_{Y_1} \mu_{Y_2}} \left( \frac{1}{\mu_{Y_1}} \sigma_{X_1 Y_1} h h + \frac{1}{\mu_{Y_2}} \sigma_{X_2 Y_1} h h \right) + \frac{2}{\mu_{Y_2} \mu_{Y_1}} \left( \frac{1}{\mu_{Y_2}} \sigma_{X_1 Y_2} h h + \frac{1}{\mu_{Y_1}} \sigma_{X_2 Y_2} h h \right) \right] \\ &+ \frac{P^2}{n} \left[ \left( \frac{1}{\mu_{Y_1}^2} \sigma_{u_1}^2 + \frac{1}{\mu_{Y_2}^2} \sigma_{u_2}^2 \right) h^2 + \left( \frac{1}{P^2 \mu_{Y_1}^2} \sigma_{v_1}^2 h^2 + \frac{1}{P^2 \mu_{Y_2}^2} \sigma_{v_2}^2 h^2 \right) \right] \end{aligned} \tag{3.3}$$

### ESTIMATOR

#### 4 PROPOSED CLASS OF GENERALIZED UNBIASED

We consider a simple random sample of size  $n=2m$  and split this sample randomly into two sub samples each of size  $m$ . Let  $(\bar{y}_{1,2m}, \bar{y}_{2,2m})$  be the product of sample means of values on  $Y_1, Y_2$  and  $(\bar{x}_{1,2m}, \bar{x}_{2,2m})$  be the sample means of  $(X_1, X_2)$  respectively

for the entire sample of size  $2m$  and  $\bar{y}_{1,m}, \bar{y}_{2,m}$  be the product of sample means of  $Y_1, Y_2$  and  $(\bar{x}_{1,2m}, \bar{x}_{2,2m})$  be the sample means of  $(X_1, X_2)$  respectively for  $i_{th}$  ( $i=1,2$ ) sub-sample of size  $m$ .

The generalized estimator based on sample of size  $2m$  of product of population means  $\mu_{Y_1} \mu_{Y_2}$  using means  $\mu_1$  and  $\mu_2$  of auxiliary variables  $X_1$  and  $X_2$ ,  $\hat{P}_H^{(i)}$  and  $\hat{P}_H^{(2)}$  be the generalized estimators for the two randomly split sub-samples of size  $m$  each, respectively, given by

$$\hat{P}_H^{(3)} = H \left( \hat{P}_{2m}, \frac{\bar{x}_{1,2m}}{\mu_1}, \frac{\bar{x}_{2,2m}}{\mu_2} \right) \tag{4.1}$$

and

$$\hat{P}_H^{(1)} = H \left( \hat{P}_m^{(1)}, \frac{\bar{x}_{1,m}^{(1)}}{\mu_1}, \frac{\bar{x}_{2,m}^{(1)}}{\mu_2} \right) \tag{4.2}$$

$$\hat{P}_H^{(2)} = H \left( \hat{P}_m^{(2)}, \frac{\bar{x}_{1,m}^{(2)}}{\mu_1}, \frac{\bar{x}_{2,m}^{(2)}}{\mu_2} \right) \tag{4.3}$$

where

$$H \left[ \hat{P}_{2m}, \frac{\bar{x}_{1,2m}}{\mu_1}, \frac{\bar{x}_{2,2m}}{\mu_2} \right], H \left[ \hat{P}_m^{(1)}, \frac{\bar{x}_{1,m}^{(1)}}{\mu_1}, \frac{\bar{x}_{2,m}^{(1)}}{\mu_2} \right] \text{ and } H \left[ \hat{P}_m^{(2)}, \frac{\bar{x}_{1,m}^{(2)}}{\mu_1}, \frac{\bar{x}_{2,m}^{(2)}}{\mu_2} \right]$$

are the bounded functions of  $\left( \hat{P}_{2m}, \frac{\bar{x}_{1,2m}}{\mu_1}, \frac{\bar{x}_{2,2m}}{\mu_2} \right)$ ,  $\left( \hat{P}_m^{(1)}, \frac{\bar{x}_{1,m}^{(1)}}{\mu_1}, \frac{\bar{x}_{2,m}^{(1)}}{\mu_2} \right)$  and  $\left( \hat{P}_m^{(2)}, \frac{\bar{x}_{1,m}^{(2)}}{\mu_1}, \frac{\bar{x}_{2,m}^{(2)}}{\mu_2} \right)$  respectively

satisfying the regularity conditions of  $h(\cdot)$  in (2.1) for  $H(\cdot)$  involved in generalized estimator  $\hat{P}_H^{(i)}$  where  $i=1, 2, 3$ . On the lines of Sukhatme and Sukhatme (Chapter IV, page 162), for  $N$  to be large we consider the Jack-knife generalized estimator  $\hat{P}_{Hj}$  given by

$$\hat{P}_{Hj} = 2\hat{P}_H^{(3)} - \frac{1}{2}(\hat{P}_H^{(1)} + \hat{P}_H^{(2)}) \tag{4.4}$$

Now, consider

$$\hat{P}_{2m} = \sum_{i=1}^{2m} y_{1,2m} \sum_{j=1}^{2m} y_{2,2m} = P \left[ 1 + \frac{W_{u_1} + W_{y_1}}{(2m)^{1/2} \mu_{Y_1}} + \frac{W_{u_2} + W_{y_2}}{(2m)^{1/2} \mu_{Y_2}} + \frac{(W_{u_1} + W_{y_1})(W_{u_2} + W_{y_2})}{(2m)^{1/2} \mu_{Y_1} (2m)^{1/2} \mu_{Y_2}} \right] \tag{4.5}$$

and  $\sum_{i=1}^m \sum_{j=1}^m$

$$\hat{P}_H^{(i)} = \sum_{i=1}^m y_{1,i} \sum_{j=1}^m y_{2,i} = P \left[ 1 + \frac{W_{u_1}^{(i)} + W_{y_1}^{(i)}}{(m)^{1/2} \mu_{Y_1}^{(i)}} + \frac{W_{u_2}^{(i)} + W_{y_2}^{(i)}}{(m)^{1/2} \mu_{Y_2}^{(i)}} + \frac{W_{u_1}^{(i)} + W_{y_1}^{(i)}}{(m)^{1/2} \mu_{Y_1}^{(i)}} \frac{W_{u_2}^{(i)} + W_{y_2}^{(i)}}{(m)^{1/2} \mu_{Y_2}^{(i)}} \right] \tag{4.6}$$

where  $i=1,2$

Also,

$$\bar{x}_{1,2m} = \frac{1}{(2m)^{1/2}} [W_{v_1} + W_{x_1}] + \mu_1 \tag{4.7}$$

$$\bar{x}_{2,2m} = \frac{1}{(2m)^{1/2}} [W_{v_2} + W_{x_2}] + \mu_2 \tag{4.8}$$

$$\bar{x}_{1,m}^{(1)} = \frac{1}{(m)^{1/2}} [W_{v_1}^{(1)} + W_{x_1}^{(1)}] + \mu_1 \tag{4.9}$$

$$\bar{x}_{2,m}^{(1)} = \frac{1}{(m)^{1/2}} [W_{v_2}^{(1)} + W_{x_2}^{(1)}] + \mu_2 \tag{4.10}$$

$$\bar{x}_{1,m}^{(2)} = \frac{1}{(m)^{1/2}} [W_{v_1}^{(2)} + W_{x_1}^{(2)}] + \mu_1 \tag{4.11}$$

$$\bar{x}_{2,m}^{(2)} = \frac{1}{(m)^{1/2}} [W_{v_2}^{(2)} + W_{x_2}^{(2)}] + \mu_2 \tag{4.12}$$

We expand (4.1), (4.2) and (4.3) similarly as (3.1) using Taylor's series expansion up to first order of approximation, to get

$$\hat{P}_H^{(3)} \cdot P = \left[ P \left\{ \frac{W_{u_1} + W_{y_1}}{(2m)^{1/2} \mu_{Y_1}} + \frac{W_{u_2} + W_{y_2}}{(2m)^{1/2} \mu_{Y_2}} + \frac{(W_{u_1} + W_{y_1})(W_{u_2} + W_{y_2})}{(2m)^{1/2} \mu_{Y_1} (2m)^{1/2} \mu_{Y_2}} \right\} H \right]_1 + \frac{W_{v_1} + W_{x_1}}{(2m)^{1/2} \mu_1} H - \frac{W_{v_2} + W_{x_2}}{(2m)^{1/2} \mu_2} H + \frac{1}{2!(2m)} \left\{ P^2 \left[ \frac{W_{u_1} + W_{y_1}}{\mu_{Y_1}} - \frac{W_{u_2} + W_{y_2}}{\mu_{Y_2}} \right]^2 H \right\}_{300} + \frac{(W_{u_1} + W_{y_1})^2}{\mu_1^2} H_{030} + \frac{(W_{u_2} + W_{y_2})^2}{\mu_2^2} H_{003} + 2 \frac{P(W_{u_1} + W_{y_1})(W_{u_2} + W_{y_2})}{\mu_{Y_1} \mu_{Y_2}} H_{110} + 2 \frac{P(W_{u_2} + W_{y_2})(W_{v_1} + W_{x_1})}{\mu_{Y_2} \mu_1} H_{110} + 2 \frac{P(W_{v_1} + W_{x_1})(W_{v_2} + W_{x_2})}{\mu_1 \mu_2} H_{011} + 2 \left[ \frac{P(W_{u_1} + W_{y_1})(W_{v_1} + W_{x_1})}{\mu_{Y_1} \mu_1} H_{101} + \frac{P(W_{u_2} + W_{y_2})(W_{v_2} + W_{x_2})}{\mu_{Y_2} \mu_2} H_{101} \right] \tag{4.13}$$

For  $i=1,2$ , we have

$$\hat{P}_H^{(i)} \cdot P = P \left[ W_{u_i} + W_{y_i} + \frac{(W_{u_i} + W_{y_i})(W_{u_i} + W_{y_i})}{(m)^{1/2} \mu_{Y_i}^{(i)}} \right] H \left[ \left\{ \frac{(i)}{m^{1/2} \mu_{Y_1}^{(i)}} - \frac{(i)}{m^{1/2} \mu_{Y_2}^{(i)}} + \frac{(i)}{m^{1/2} \mu_{Y_1}^{(i)}} \frac{(i)}{m^{1/2} \mu_{Y_2}^{(i)}} \right\} \right]_1 + \frac{W_{v_1}^{(i)} + W_{x_1}^{(i)}}{m^{1/2} \mu_1^{(i)}} H_2 + \frac{W_{v_2}^{(i)} + W_{x_2}^{(i)}}{m^{1/2} \mu_2^{(i)}} H_3 + 2! m \left\{ P^2 \left[ \frac{W_{u_1}^{(i)} + W_{y_1}^{(i)}}{\mu_{Y_1}^{(i)}} - \frac{W_{u_2}^{(i)} + W_{y_2}^{(i)}}{\mu_{Y_2}^{(i)}} \right]^2 H \right\}_{300} + \frac{(W_{v_1}^{(i)} + W_{x_1}^{(i)})^2}{\mu_1^2} H_{030} + \frac{(W_{v_2}^{(i)} + W_{x_2}^{(i)})^2}{\mu_2^2} H_{003} + 2 \frac{P(W_{u_1}^{(i)} + W_{y_1}^{(i)})(W_{v_1}^{(i)} + W_{x_1}^{(i)})}{\mu_{Y_1}^{(i)} \mu_1^{(i)}} H_{110}$$

$$\begin{aligned}
 &+2 \frac{P(W_{u_1}^{(i)} + W_{y_1}^{(i)})(W_{v_1}^{(i)} + W_{x_1}^{(i)})}{\mu_{y_1} \mu_{x_1}} H_{110} + 2 \frac{P(W_{u_2}^{(i)} + W_{y_2}^{(i)})(W_{v_2}^{(i)} + W_{x_2}^{(i)})}{\mu_{y_2} \mu_{x_2}} H_{011} \\
 &+ 2 \left. \left. \frac{P(W_{u_1}^{(i)} + W_{y_1}^{(i)})(W_{v_1}^{(i)} + W_{x_1}^{(i)})}{\mu_{y_1} \mu_{x_1}} H_{101} + 2 \frac{P(W_{u_2}^{(i)} + W_{y_2}^{(i)})(W_{v_2}^{(i)} + W_{x_2}^{(i)})}{\mu_{y_2} \mu_{x_2}} H_{101} \right\} P_{101} \right\} \quad (4.14)
 \end{aligned}$$

Substituting these values in (4.4), we have

$$\begin{aligned}
 \hat{P}_{Hj} &= \frac{1}{P} \left[ \frac{W_u + W_y}{(2m)^{1/2} \mu_{y_1}} + \frac{W_v + W_x}{(2m)^{1/2} \mu_{y_2}} + \frac{(W_u + W_y)(W_v + W_x)}{(2m)^{1/2} \mu_{y_1} (2m)^{1/2} \mu_{y_2}} \right] H \\
 &+ \frac{W_{v_1} + W_{x_1}}{(2m)^{1/2} \mu_{x_1}} H_2 + \frac{W_{v_2} + W_{x_2}}{(2m)^{1/2} \mu_{x_2}} H_3 \\
 &+ \frac{1}{2!(2m)} \left\{ P^2 \left[ \frac{W_{u_1} + W_{y_1}}{\mu_{y_1}} + \frac{W_{u_2} + W_{y_2}}{\mu_{y_2}} \right]^2 \right\} H_{300} \\
 &+ \frac{W_{v_1} + W_{x_1}}{\mu_{x_1}^2} H_{030} + \frac{W_{v_2} + W_{x_2}}{\mu_{x_2}^2} H_{003} \\
 &+ 2 \frac{(W_{v_1} + W_{x_1})(W_{v_2} + W_{x_2})}{\mu_{y_1} \mu_{y_2}} H_{011} \\
 &+ 2 \frac{P(W_{u_1} + W_{y_1})(W_{v_1} + W_{x_1})}{\mu_{y_1} \mu_{x_1}} H_{110} \\
 &+ 2 \frac{H(W_{u_2} + W_{y_2})(W_{v_1} + W_{x_1})}{\mu_{y_2} \mu_{x_1}} H_{110} \\
 &+ 2 \frac{P(W_{u_1} + W_{y_1})(W_{v_2} + W_{x_2})}{\mu_{y_1} \mu_{x_2}} H_{101} \\
 &+ 2 \frac{P(W_{u_2} + W_{y_2})(W_{v_2} + W_{x_2})}{\mu_{y_2} \mu_{x_2}} H_{101} \left. \right\} \\
 &- \frac{1}{2!} \sum_{i=1}^2 \left[ \frac{P(W_{u_1}^{(i)} + W_{y_1}^{(i)})}{\mu_{y_1}} + \frac{P(W_{u_2}^{(i)} + W_{y_2}^{(i)})}{\mu_{y_2}} \right. \\
 &\left. \frac{1}{(2m)} \left[ \frac{W_{u_1} + W_{y_1}}{\mu_{y_1}} + \frac{W_{u_2} + W_{y_2}}{\mu_{y_2}} \right] \right] H_1 \\
 &+ \frac{W_{v_1}^{(i)} + W_{x_1}^{(i)}}{m^{1/2} \mu_{x_1}} H_2 + \frac{W_{v_2}^{(i)} + W_{x_2}^{(i)}}{m^{1/2} \mu_{x_2}} H_3 \\
 &+ \frac{1}{2!m} \left\{ P^2 \left[ \frac{W_{u_1} + W_{y_1}}{\mu_{y_1}} + \frac{W_{u_2} + W_{y_2}}{\mu_{y_2}} \right]^2 \right\} H_{300} \\
 &+ \frac{(W_{v_1}^{(i)} + W_{x_1}^{(i)})^2}{\mu_{x_1}^2} H_{030} + \frac{(W_{v_2}^{(i)} + W_{x_2}^{(i)})^2}{\mu_{x_2}^2} H_{003} \\
 &+ 2 \frac{P(W_{u_1} + W_{y_1})(W_{v_1}^{(i)} + W_{x_1}^{(i)})}{\mu_{y_1} \mu_{x_1}} H_{110} \\
 &+ 2 \frac{P(W_{u_2}^{(i)} + W_{y_2}^{(i)})(W_{v_1}^{(i)} + W_{x_1}^{(i)})}{\mu_{y_2} \mu_{x_1}} H_{110}
 \end{aligned}$$

$$\begin{aligned}
 &+ 2 \frac{P(W_{v_1} + W_{x_1})(W_{v_2} + W_{x_2})}{\mu_{y_1} \mu_{y_2}} H_{011} \\
 &+ 2 \frac{P(W_{u_1} + W_{y_1})(W_{v_2} + W_{x_2})}{\mu_{y_1} \mu_{x_2}} H_{101} \\
 &+ 2 \frac{P(W_{u_2} + W_{y_2})(W_{v_2} + W_{x_2})}{\mu_{y_2} \mu_{x_2}} H_{101} \left. \right\} \quad (4.15)
 \end{aligned}$$

On taking expectation on both sides (4.15) the bias of the  $\hat{P}_{Hj}$

up to terms of order  $O\left(\frac{1}{m}\right)$ , is given by

$$\text{Bias } \hat{P}_{Hj} = O\left(\frac{1}{m}\right) \quad (4.16)$$

On squaring both side of (4.15) and taking expectation, the

mean square error of  $\hat{R}_{Gj}$  up to terms of order

$O\left(\frac{1}{m}\right)$ , is given by

$$\begin{aligned}
 \text{MSE } \hat{P}_{Hj} &= E \left[ \hat{P}_{Hj} - P_2 \right]^2 \\
 &= E \left[ 2P \left[ \frac{W_{u_1} + W_{y_1}}{(2m) \mu_{y_1}} + \frac{W_{u_2} + W_{y_2}}{(2m) \mu_{y_2}} \right] H_1 + \frac{W_{v_1} + W_{x_1}}{(2m) \mu_{x_1}} H_2 \right. \\
 &\quad + \frac{W_{v_2} + W_{x_2}}{(2m) \mu_{x_2}} H_3 - \frac{1}{2} \sum_{i=1}^2 \left[ \frac{W_{u_1}^{(i)} + W_{y_1}^{(i)}}{m^{1/2} \mu_{y_1}} + \frac{W_{u_2}^{(i)} + W_{y_2}^{(i)}}{m^{1/2} \mu_{y_2}} \right] \\
 &\quad \left. + \frac{W_{v_1}^{(i)} + W_{x_1}^{(i)}}{m^{1/2} \mu_{x_1}} H_2 + \frac{W_{v_2}^{(i)} + W_{x_2}^{(i)}}{m^{1/2} \mu_{x_2}} H_3 \right]^2 \right] \\
 &= 2m \left[ \frac{1}{\mu_{y_1}^2} \sigma_{y_1}^2 + \frac{1}{\mu_{y_2}^2} \sigma_{y_2}^2 + \frac{1}{\mu_{x_1} \mu_{x_2}} \sigma_{y_1 y_2} \right] H_1^2 \\
 &\quad + \left( \frac{1}{P^2 \mu_{x_1}^2} \sigma_{x_1}^2 H_2^2 + \frac{1}{P^2 \mu_{x_2}^2} \sigma_{x_2}^2 H_3^2 + \frac{2}{P^2 \mu_{x_1} \mu_{x_2}} \sigma_{x_1 x_2} H_2 H_3 \right) \\
 &\quad + \frac{2}{\mu_{y_1}} \left( \frac{1}{\mu_{x_1}} \sigma_{x_1 y_1} H_1 H_2 + \frac{1}{\mu_{x_2}} \sigma_{x_2 y_1} H_1 H_3 \right) \\
 &\quad + \frac{2}{\mu_{y_2}} \left( \frac{1}{\mu_{x_1}} \sigma_{x_1 y_2} H_1 H_2 + \frac{1}{\mu_{x_2}} \sigma_{x_2 y_2} H_1 H_3 \right) \\
 &\quad + \frac{P}{2m} \left[ \frac{1}{\mu_{y_1}^2} \sigma_{y_1}^2 + \frac{1}{\mu_{y_2}^2} \sigma_{y_2}^2 \right] H^2 \\
 &\quad + \left( \frac{1}{P^2 \mu_{x_1}^2} \sigma_{x_1}^2 H_2^2 + \frac{1}{P^2 \mu_{x_2}^2} \sigma_{x_2}^2 H_3^2 \right) \quad (4.17)
 \end{aligned}$$

### 5 CONCLUDING REMARK

For the proposed generalized estimators  $\hat{P}_h$  and  $\hat{P}_{Hj}$ , it can be easily seen from (3.2), (3.3), (4.16) and (4.17) that bias and the

MSE are given by

$$\text{Bias } \left( \hat{P}_h \right) = \frac{1}{\pi} \left[ \frac{1}{2\mu_{x_1}^2} \sigma_{x_1}^2 h + \frac{1}{2\mu_{x_2}^2} \sigma_{x_2}^2 h + \frac{1}{\mu_{x_1} \mu_{x_2}} \sigma_{x_1 x_2} h \right] \quad (8, 9)$$

$$\begin{aligned}
 & + \frac{1}{\mu} \left[ \frac{1}{\mu} \sigma_{(x_1, y_1)} h_{110} + \frac{1}{\mu} \sigma_{(x_2, y_1)} h_{101} \right] \\
 & + \frac{1}{\mu} \left[ \frac{1}{\mu} \sigma_{(x_1, y_2)} h_{110} + \frac{1}{\mu} \sigma_{(x_2, y_2)} h_{101} \right] \\
 & + \frac{1}{n} \left[ \frac{1}{2\mu^2} \sigma_{v_1}^2 h_{030} + \frac{1}{2\mu^2} \sigma_{v_2}^2 h_{003} \right] \quad (5.1)
 \end{aligned}$$

measurement error as they help in removing the bias while still preserving the efficiency. Many important estimators like adapted estimators of Srivastava (1967) type for estimation of ratio of population means given by

$$\hat{P}_\alpha = \hat{P} \left( \frac{\bar{x}_1}{\mu_1} \right)^{\alpha_1} \left( \frac{\bar{x}_2}{\mu_2} \right)^{\alpha_2}$$

are special cases of the present study when we assume the measurement errors to be absent. Such conventional results can be obtained as special cases of this study by setting the

measurement error variances to be zero or by choosing the derivatives judiciously under measurement errors.

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$$\begin{aligned}
 \text{MSE}(\hat{P}_h) &= \frac{P^2}{n} \left[ \left( \frac{1}{\mu^2} \sigma_{y_1}^2 + \frac{1}{\mu^2} \sigma_{y_2}^2 + \frac{1}{\mu} \sigma_{(y_1, y_2)} \right) h^2 \right. \\
 & + \left. \left( \frac{1}{P^2 \mu_1^2} \sigma_{x_1}^2 + \frac{1}{P^2 \mu_2^2} \sigma_{x_2}^2 + \frac{2}{P \mu_1 \mu_2} \sigma_{(x_1, x_2)} \right) h^2 \right] \\
 & + \frac{2}{n} \left[ \frac{1}{\mu} \sigma_{(x_1, y_1)} h h + \frac{1}{\mu} \sigma_{(x_2, y_1)} h h \right] \\
 & + \frac{2}{n} \left[ \frac{1}{\mu} \sigma_{(x_1, y_2)} h h + \frac{1}{\mu} \sigma_{(x_2, y_2)} h h \right] \\
 & + \frac{1}{n} \left[ \left( \frac{1}{\mu^2} \sigma_{v_1}^2 + \frac{1}{\mu^2} \sigma_{v_2}^2 \right) h^2 \right] \\
 & + \left( \frac{1}{P^2 \mu_1^2} \sigma_{x_1}^2 + \frac{1}{P^2 \mu_2^2} \sigma_{x_2}^2 \right) \quad (5.2)
 \end{aligned}$$

$$\text{Bias}(\hat{P}_{Hj}) = 0 \quad (5.3)$$

$$\begin{aligned}
 \text{MSE}(\hat{P}_{Hj}) &= \frac{P_2}{2m} \left[ \left( \frac{1}{\mu^2} \sigma_{y_1}^2 + \frac{1}{\mu^2} \sigma_{y_2}^2 + \frac{1}{\mu} \sigma_{(y_1, y_2)} \right) H^2 \right. \\
 & + \left. \left( \frac{1}{P^2 \mu_1^2} \sigma_{x_1}^2 + \frac{1}{P^2 \mu_2^2} \sigma_{x_2}^2 + \frac{2}{P \mu_1 \mu_2} \sigma_{(x_1, x_2)} \right) H^2 \right] \\
 & + \frac{2}{m} \left[ \frac{1}{\mu} \sigma_{(x_1, y_1)} H H + \frac{1}{\mu} \sigma_{(x_2, y_1)} H H \right] \\
 & + \frac{2}{m} \left[ \frac{1}{\mu} \sigma_{(x_1, y_2)} H H + \frac{1}{\mu} \sigma_{(x_2, y_2)} H H \right] \\
 & + \frac{1}{2m} \left[ \left( \frac{1}{\mu^2} \sigma_{v_1}^2 + \frac{1}{\mu^2} \sigma_{v_2}^2 \right) H^2 \right] \\
 & + \left( \frac{1}{P^2 \mu_1^2} \sigma_{x_1}^2 + \frac{1}{P^2 \mu_2^2} \sigma_{x_2}^2 \right) \quad (5.4)
 \end{aligned}$$

From (5.2) and (5.4), we see that both estimators  $\hat{P}_h$  and  $\hat{P}_{Hj}$  have the same mean square error but from (5.1) bias of  $\hat{P}_h$  is not zero whereas from (5.3) bias of the jack-knifed estimator  $\hat{P}_{Hj}$  is zero. Hence, both the estimators  $\hat{P}_h$  and  $\hat{P}_{Hj}$  having the same mean square error, but the jack-knifed estimator  $\hat{P}_{Hj}$  may be preferred to the estimator  $\hat{P}_h$  in the sense of unbiasedness under the presence of measurement errors. Therefore, the proposed unbiased jack-knifed estimator should be preferred than the conventional estimators under